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To measure or not to measure?
An empirical investigation of social impact measurement in UK social enterprises

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Abstract

Social enterprises (SE) – organisations with a dual mission to generate economic and social value – have become important players in the delivery of public services in the UK and elsewhere. While public sector value-for-money imperatives encourages these hybrid organisations to provide estimates of their social and economic impact, relatively little is known about who does so. Using institutional perspectives and large-sample data produced by Social Enterprises UK, we empirically document the uptake of social impact measurement in this sector and the extent to which context, the nature of the impact, and stakeholders involvement explain differences in participation rates.

Keywords: Social enterprises; social impact measurement; government funding; institutional pressures; United Kingdom.
Social enterprises (SEs) pursue social missions using market mechanisms. As such they face complex governance challenges. They produce both public and private goods, and since the value of the former is typically much more difficult to measure, the tendency is to emphasize the latter. This is consistent with dominant traditional accounting paradigms that emphasize financial results and easily measured outcomes (Ebrahim, Battilana and Mair 2014; Gibbon and Dey 2011; Millar and Hall 2013).

This tendency has intensified with the move towards New Public Management (NPM) promoted in Europe and elsewhere as resources-constrained welfare systems increasingly rely on SEs and non-profit organisations to design and deliver public services (Powell, Gillett and Doherty 2018; Doherty, Haugh and Lyon 2014; Arvidson and Lyon 2014; Mason 2012).

NPM offers prescriptions, by which public sector organisations should be designed, organized, managed and, ultimately, function. The overarching principle of NPM is to make public sector organisations, and by default those interacting with it, ‘performance, cost, efficiency, and audit-oriented’ (Diefenbach 2009: 893). Consequently, NPM-inspired social impact measurement schemes, such as Social Return on Investment (SROI) and Social Accounting and Audit (SAA), have become an important part of governmental regulatory approval processes in the assignment of contracts, grants and other things to social enterprises (Fazzi 2012; Cunningham, Baines and Charlesworth 2014; O’Dwyer and Unerman 2007; Millar and Hall 2013).

For many proponents, social impact measurement/assessment is more than a stand-alone process limited to the measurement of social outputs. Full engagement with these approaches can help SEs run more effectively and keep operations aligned to missions (e.g. McLoughlin, Kaminski, Sodagar, Khan, Harris, Arnaudo and McBreary 2009; Barman 2007). Moreover, SEs that embrace social impact measurement may be more likely to create the participatory and
deliberative processes that facilitate community discussions about the proposed social impacts of
the organisation and its activities (Esteves, Franks and Vanclay 2012; Fazzi 2012; Arvidson and
Lyon 2014). Social impact measurement can also act as an organisational legitimacy control
verifying that the SE has respected its self-imposed rules (statute, mission, program of action) and
the legal norms applicable to its institutional formula (Suchman 1995; Nicholls 10b; Bagnoli and
Megali 2011).

In practice however, SEs pursue very different social missions which makes comparisons
of social impact complex and challenging (Doherty et al 2014; Millar and Hall 2013). In other
words, the concept of social impact lacks coherence and robustness, and since it is largely self-
determined, it can be subject to manipulation (Arvidson and Lyon 2014; Barman 2007). In
particular, critics argue that NPM approaches to social impact measurement promote one-
dimensional focus on funder perspectives, invariably prioritizing investing stakeholders over
others (Defourny and Nyssens 2010; Esteves et al 2012; Julnes and Holzer 2001; Millar and Hall
2013). This, they claim, encourages mission drift towards the objectives of outside resource
providers and amplifies the risk of managerial capture and political hijacking (e.g. Diefenbach
2009; Ebrahim et al 2014; Powell and Osborne 2020). Also, when accountability schemes are
perceived as being controlled by ‘outside’ stakeholders for purposes of comparisons with
competitors and/or to oversee performance management, they can have detrimental effects on SE
organisational culture and staff morale (Christensen and Ebrahim 2006; Hwang and Powell 2009;
Ebrahim 2005; Gibbon and Dey 2011).

Accordingly, and as the social enterprise model continues to spread, so is the realization
that social accounting frameworks are not only inadequate for this hybrid organisational form, but
are also damaging its development and future sustainability (O’Dwyer and Unerman 2007; Liston-
Heyes, Hall, Jevtovic and Elson 2017). These tensions suggest that there may be obstacles to the adoption of social impact measurement schemes beyond awareness and resource constraints.

Given the scope of the debate, it is surprising how little is known about social measurement by SEs in practice. Exceptions include Bertotti, Leahy, Sheridan, Tobi and Renton (2011) and Leah and Villeneuve-Smith (2009) who provide descriptive statistics focused on the UK health and social care sectors using the State of Social Enterprise Survey 2009. Another is a study by Maas and Grieco (2017) who use an international sample of 3194 SEs from the Global Entrepreneurship Monitor data to investigate the relationship between the nature of SE mission and its decision to measure social impact. Our analysis complements and extends this research by investigating the factors that motivate (or hinder) social impact measurement in UK-based SEs that participated in the 2017 State of Social Enterprise Survey. As in earlier studies, the data we use is self-reported and subject to variation in what constitutes social impact and how it is measured.

Informed by the literature and guided by the concept of organisational legitimacy, we present a model that connects the measurement of social impact to SE attributes. More concretely, we posit that the SE context, the nature of its impact, non-funding and funding stakeholders’ engagement in SE decisions will influence propensities to measure social impact. This theoretical discussion appears in the first section of the paper. Section 2 explains the data and the ordered and simple Logit regression approaches we use to test the hypotheses. Results are presented in Section 3, and discussed in Section 4. The paper ends with brief conclusions highlighting contributions to the academic and practitioner literatures and important caveats.
Social enterprises and social impact measurement

We use institutional theories to conceptualise the pressures that frame SEs’ decisions to measure their social impact and in the development of testable hypotheses. Our model is informed by DiMaggio and Powell (1983)’s isomorphic processes (coercive, mimetic, and normative) which provide useful tools in identifying the forces that regularise the sector and homogenise its practices. We also invoke Suchman’s (1995) who revisits these dynamics through the lens of organisational legitimacy. Organisational legitimacy incorporates both institutional legitimacy – which focuses on the pressures and dynamics that transcend any single organisation’s purposive control – and strategic legitimacy – which emphasises the managerial perspective and instrumental manipulations by organisations to garner societal support. As with other SE researchers, we argue that these perspectives are useful in describing the events that shaped the development of the sector and in explaining SEs’ stance towards social impact measurement (e.g. Arvidson and Lyon 2014; Barman 2007; Nicholls 2010b; Bagnoli and Megali 2011). Evidence suggests that for UK SEs, organisational legitimacy has and continues to be driven by public sector ideologies, norms, and regulations (Teasdale, Alcock and Smith 2012; Powell et al 2019).

The perceived potential of SEs to operate as surrogate public organisations was highlighted by the severe austerity measures that followed the financial crisis, pressuring governments to do ‘more with less’. Framed by notions of ‘Big Society’ and ‘Third Way’, the SE sector experienced unprecedented growth as governments reduced their direct involvement in public service delivery while encouraging SEs to fill this gap through government grants and contracts (Power and

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3 The analyses that follows are based on the frequently cited definition by the UK Department of Trade and Industry which refers to an SE as ‘(…) a business with primarily social objectives whose surpluses are principally reinvested for that purpose in the business or in the community, rather than being driven by the need to maximize profit for shareholders and owners’ (SE Market Trends 2017, 14). (For discussions of SE definitions, see Defourny and Nyssens (2010), Powell, Gillett, and Doherty (2019), and Doherty, Haugh, and Lyon (2014)).
Osborne 2020; Dey and Teasdale 2016; Hall, Miller, and Millar 2016). Concerns emerged as to whether the potential of SEs in public service delivery had been exaggerated and whether this hybrid form was financially viable in the long term (e.g. Mason 2012; Powell et al 2019; VanSandt et al 2009; Sud, VanSandt and Baugous 2009). Perhaps unsurprisingly, the expansion of the sector was accompanied by pressures to demonstrate its achievements through NPM-inspired performance measurement and value-for-money principles (Millar and Hall 2013; Barman 2007). For SEs, failure to do so could reduce the visibility of their contribution and undermine their access to resources (funds and in-kind). Moreover, for policy-makers demonstrating such value could help legitimise the transfer of social responsibilities and accelerated growth of the sector to the wider public (Mason 2012; Nicholls 2010b; Gibbon and Dey 2011).

Responses to institutional pressures to measure performance differs across the SE sector. Some welcome social measurement schemes as tools that help demonstrate and ‘frame’ SE effectiveness to external stakeholders, providing a competitive advantage in the tender of public sector contracts and grant applications (Ryan and Lyne 2008; Peattie and Morley 2008; Lee and Huang 2018). Such schemes can also serve a pedagogic function by providing guidance and control for the organisation and helping staff to analyse and improve their services (Arvidson and Lyon 2014; Diefenbach 2009; Vansandt et al 2009). On the other hand, for some SEs measuring social impact can be prohibitively costly and/or divert too many resources away from key activities. This is particularly so when most of a SE’s activities are ‘soft’ and require subjective value judgements (Millar and Hall 2013; Kendall and Knapp 2000). These valuations can also become contentious when they are used for competitive comparisons in the allocation of funding (Ryan and Lyne 2008; Christensen and Ebrahim 2006; Hwang and Powell 2009; Ebrahim 2005). Social impact measures are also associated with managerial capture, mission drift, and in some
cases, tensions linked to ideological and cultural differences of opinions with respect to the necessity and organisational consequences of measuring social value (Millar and Hall 2013; Diefenbach 2009; Cunningham et al 2014). Developing and using performance measures of any kind often involves transformations that may be threatening to an organisation - whether these threats are real or not - and lead to long term decreases in actual performance (Julnes and Holzer 2001; Arvidson and Lyon 2014; Diefenbach 2009).

In other words, there are substantial pressures for SEs to mimic the management systems of public sector agencies but there are also grounds for resistance. This may explain, at least in part, the relatively slow uptake of social measurement tools in the SE sector (Peattie and Morley 2008; Bertotti et al 2011; Millar and Hall 2013). Against this background, we investigate some of the institutional factors that may be encouraging/discouraging social impact measurement in SEs.

While our analysis focuses on the dichotomous ‘to measure or not to measure’ decision, we recognize that social impact is socially constructed and there are tensions over how it is defined and measured (Barman 2007; Hwang and Powell 2009). These tensions can incentivize impression management, decoupling, symbolic compliance, and deflection amongst other strategic and/or coping behaviours engendered by compliance with the performance measurement process (Julnes and Holzer 2001; Dey and Teasdale 2016; Arvidson and Lyon 2014). These are important and connected issues but we lack the data for a more nuanced investigation. Nonetheless, we argue that a systematic inquiry into the patterns of social impact measurement adoption can provide preliminary insights into SE pressures for conformance to sector norms (Ebrahim and Rangan 2014). Implicitly, we recognize that the context in which the more general decision to measure

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4 As noted by one reviewer, a more thorough investigation would also examine the extent to which public servants use SEs’ social impact measures to demonstrate the legitimacy of their actions and progress their own agenda.
social impact is as important, if not more, than the actual measures that are used (Julnes and Holzer 2001).

With this in mind, we propose four sets of factors reflecting different isomorphic pressures that influence the SE’s decision to measure social impact (Nicholls 2010a; Mason 2012; Dey and Teasdale 2016; Arvidson and Lyon 2014).

\textit{HI – SE Context}

Measuring social impact is a knowledge-intensive endeavour (Julnes and Holzer 2001; Kendall and Knapp 2000; Liston-Heyes et al 2017). Accordingly, we posit that SEs operating in environments where this knowledge is more easily accessed will find it easier to measure their social impact (Millar and Hall 2013).

One implication is that SEs with the knowledge to measure impact are more likely to be located in areas with greater social capital. More generally, geographical proximity or co-presence to wealthier urban centres facilitates philanthropic support of all kind including volunteering, donations (monetary and in-kind), participation to advisory board membership, and access to corporate resources (Liston-Heyes et al 2017; Mason 2012; McCulloch et al 2012; Mohan 2012). SEs that are embedded in social networks will have easier access to the ‘new breed’ of paid managers that have been trained in the art of NPM and value-for-money principles (Diefenbach 2009; DiMaggio and Powell 1983; Ebrahim and Rangan 2014; Millar and Hall 2013). These employees tend to have more homogeneous preferences aligned with greater formalization and NPM ideals (Fazzi 2012). In other words, isomorphic pressures on SEs will vary considerably according to geography (Clifford 2012; Mohan 2012). In the UK, one expects higher levels and/or easier access to social capital, greater levels of professionalization, and the ratification of common
administrative norms in more affluent areas that are only available to those with the resources to locate there (McCulloch et al 2012; Ebrahim and Rangan 2014; Esteves, Franks and Vanclay 2012). For these reasons, we posit that SEs located in London will have easier access to resources that facilitate social impact measurement than those based elsewhere (H1a).

Other than location, we postulate that social franchises may find it easier and/or be under greater pressure to measure their social impact. Social franchising involves a SE (the franchisor) licensing its business operating systems, products/services, or branding to other SEs (the franchisees) in exchange for agreed fees or sale-based payment (Lyon and Fernandez 2012; Tracey and Jarvis 2007). Adoption of the social franchising format can thus provide SEs with additional conformance pressures and an organisational format that has already been tested for financial viability and social impact (Bloom and Chatterji 2009; Dees, Anderson, and Wei-Skillern 2004; DiMaggio and Powell 1983; VanSandt, Sud, and Marmé 2009). Ebrahim and Rangan (2014) argue that organisational growth, coalition building and replication can facilitate legitimization and the survival of the organisation. Accordingly, we posit that SEs who have replicated or franchised their operations will have easier access to the knowledge and systems that facilitate social impact measurement (H1b).

We also hypothesize that SEs that self-describe as ‘COOPs’ will be relatively more resistant to isomorphic pressures than their counterparts. SE who perceive themselves as such will be more anchored in the cooperative traditions of collective social action and this will impact on accountability responsibilities and preferences (Doherty et al 2014; Defourny and Nyssens 2010; Ebrahim et al 2014; Powell et al 2019). These organisations run themselves along collectivist-democratic lines and pride themselves on being self-managed without outside interference (Rothschild 2009). COOPs promote community-based structures, which increases the involvement
of members without professional management skills in the governance of the SE. This can dilute direct and indirect pressures from external stakeholders to conform to sector norms (Cornforth 2014). Gibbon and Dey (2011) suggests that these organisations are particularly resistant to proposals that privilege financial over social imperatives. Additionally, in many countries, SEs predominantly assume the form of the cooperative enterprise (Fazzi 2012). For these reasons, we posit that SEs who identify themselves as COOPs are less likely to engage with social impact measurement schemes. The following hypotheses reflect these conjectures:

\[ H1a:\ \text{SEs that are based in London are more likely to measure social impact.} \]
\[ H1b:\ \text{SEs that are franchised are more likely to measure social impact.} \]
\[ H1c:\ \text{SEs that identify themselves as ‘COOPs’ are less likely to measure social impact.} \]

\textit{H2 – Nature of SE impact}

By definition, SEs will be involved in trading activities, or ‘earned income’, of one kind or another in the pursuit of their social objectives.\(^5\) They will, however, differ in how they perceive the impact they are having on society. More concretely, some SEs consider the direct provision of social, community, and environmental services as their main social contribution (Lee and Huang 2018). Others believe that their social impact is realised through the employment of disadvantaged people. There are also SEs whose focus is not on the direct provision of social goods, but rather on generating revenues for parent or partner organisations involved in solving specific social problems. Impact measurement in such cases can be potentially easier to produce (Bagnoli and Megali 2011). In other words, the challenges of measurement will differ within the SE sector

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\(^5\) Earned income is revenue generated from the sale of goods, services rendered, processes, expertise and intellectual property or work performed. This includes membership, user, program, admission, rental, and/or performance fees, conferences, symposia, event, and/or presentation services, sale of goods (new and/or used), tuition, training materials, food and catering services, newsletters, magazines, advertising sales, information products, consulting services etc. (Pue 2019).
according to the nature of the social impact created by the organisation (Doherty et al 2014; Millar and Hall 2013; Bagnoli and Megali 2011). Ebrahim and Rangan (2014) also highlight the difficulties of measuring social outcomes in SEs where the full value of social contributions is moderated by events beyond organisational boundaries. They also suggest that SEs with narrower scope of activities and shorter time lines will find it easier, less costly, and less controversial to develop and present performance indicators. Boyne (2002) and Hall et al (2016) argue that organisations exhibit different levels of ‘publicness’, and SEs that are closer to the private sector model (i.e. profit maximizing profits for shareholders and owners) will tend to have less ambiguous, and more measurable goals.

Given differences in the ‘measurability’ of SEs’ impact, we posit that measurement decisions will also depend on the nature of the social impact SEs believe they are producing (Maas and Grieco 2017; Kendall and Knapp 2000; Millar and Hall 2013). This is captured in the following:

\( H2a: \) SEs that create impact mainly through the delivery of public services are more likely to measure social impact.
\( H2b: \) SEs that create impact mainly through who they employ are less likely to measure social impact.
\( H2c: \) SEs that create impact mainly through the profits they create or gift to other social good producers are more likely to measure social impact.

\( H3 – Non-Funding Stakeholder Involvement in SE Decisions \)

Our third set of hypotheses suggests that SEs’ willingness to measure social impact will depend on the extent to which its stakeholders are involved in the process. More concretely, we recognize that SEs are pressured by funders to measure impact (see H4) but they may also do so to achieve organizational legitimacy vis-à-vis stakeholders and manage their demands on the
organisation (Esteves et al 2012; Arvidson and Lyon 2014; Diefenbach 2009; Kendall and Knapp 2000). According to Fazzi (2012), a multi-stakeholder decision-making process is the best guarantee of transparency and control. Stakeholder-inclusive governance models facilitate and enhance institutional and strategic legitimacy (Bagnoli and Megali 2011; Nicholls 2010a; Vansandt et al 2009; Esteves et al 2012), promote value co-creation and generate a competitive advantage with target communities (Powell, Gillett and Doherty 2019; Osborne, Radnor and Strokosch 2016). Lee and Huang (2018) also explore how careful framing through social impact measures can mobilise support from different stakeholders and facilitate the realization of organisational goals. On the other hand, failure to engage staff and other stakeholders in the process can lead to emasculation and feelings that the scheme is ‘controlling’ rather than ‘supportive’ (Gibbon and Dey 2011; Boyne 2002; Fazzi 2012) and increase stakeholder-related risks to the organisation (Esteves et al 2012). In other words, the extent to which internal and external stakeholder groups such as staff, consumers/beneficiaries, and the wider community are involved in the SE’s efforts at self-evaluation can facilitate and motivate the adoption of measurement schemes (Bagnoli and Megali 2011; Julnes and Holzer 2001; Arvidson and Lyon 2014). Accordingly, we propose the following:

\[ H3a: \] SEs that involve their community in their decision-making processes are more likely to measure social impact.
\[ H3b: \] SEs that involve their staff in their decision-making processes are more likely to measure social impact.
\[ H3c: \] SEs that involve their beneficiaries in their decision-making processes are more likely to measure social impact.

\[ H4 – Funding Stakeholders \]
As discussed above, the hybrid SE form grew in popularity when the UK endorsed ‘Third Way’ and ‘Big Society’ ideals that emphasised a heightened role for volunteering and social entrepreneurship in the delivery of public services (Doherty et al 2014; Millar and Hall 2013; Powell et al 2019; Hall et al 2016). Its growth accelerated further when the conservative government implemented massive public expenditure reductions in an effort to contain the economic ramifications of the 2007 financial crisis (Cunningham et al 2014; Mohan 2012). While delegating some of its responsibilities for the production of social goods by outsourcing of contracts and the allocation of grants, public sector managers retained considerable power on these organisations through NPM-inspired reporting and auditing requirements (Gibbon and Dey 2011; Powell 2007; Kendall and Knapp 2000; Fazzi 2012). Social impact indicators can also provide useful legitimation tools to policy-makers in response to critics of public sector outsourcing and high expectations of the electorate (Mason 2012; Nicholls 2010b; Gibbon and Dey 2011).

Arvidson and Lyon (2014) report SEs observing marked changes in the frequency with which public sector grant bodies referred to social impact measures and evaluation tools. These insights reflect the scale and scope of the isomorphic pressures on SEs to conform to performance measurement norms exerted by public sector managers operating under value-for-money and transparency mantras. In so far as NPM emulates techniques from the private sector and that SEs also receive funding from private sector organisations (Ebrahim and Rangan 2014; Boyne 2002; Liston-Heyes et al 2017; McCulloch et al 2012; Sud et al 2009), we hypothesise that reliance on government funding and trade with the private sector (as opposed to reliance on trade with the general public, other third sector organisations, membership fees, donations and other forms of revenue collection) will incentivise the adoption of social measurement schemes:
H4a: SEs whose main source of income is from the public sector are more likely to measure social impact.
H4b: SEs whose main source of income is from the private sector are more likely to measure social impact.

Controls

Many factors other than those hypothesized above are likely to influence the SE’s decision to measure social impact. In particular, a SE’s financial resources can facilitate the measurement process and size can ensure that the SE reaps any economies of scale and scope associated with the implementation of a measurement system (Julnes and Holzer 2001; Kendall and Knapp 2000). While we do not have access to reliable SE financial data, we used two indicators of ‘size’ to capture these effects (i.e. paid employees, and scale of operation).

Another key control routinely used in entrepreneurship studies is age of the organisation. Since younger SEs have limited resources, they are more likely to seek start-up funding from foundations and granting agencies with stricter social reporting norms (Liston-Heyes et al 2017). Relatedly, the strategic benefits of ‘venture framing’ through social impact measurement are much more consequential in the early-stage of an organisation as funders and other stakeholders have little else on which to base their assessments (Lee and Huang 2018). To alleviate the risks inherent in the management of these new business forms, younger entrepreneurs are more likely to mimic the dominant models promoted by their peers (VanSandt et al 2009; DiMaggio and Powell 1983). Accordingly, we control for SE age, SE size, and SE operating scale throughout our analyses.

Sources and Methods
We obtained data from the State of Social Enterprise Survey 2017 commissioned by Social Enterprise UK. Survey responses were gathered via telephone interviews and online surveys primarily with the individual in day-to-day control of the business and/or the individual specifically responsible for the financial affairs of the SE. A two-step filter was used to ensure that the sample reflected the landscape of social enterprises, i.e. organisations were only considered to be in the scope of the survey if they defined themselves as a social enterprise and generated 25% or more of their income from trading activities.

This produced a sample of 1060 useable responses, the largest dataset on the state of social enterprises in the UK available at the time. While the field research team (BMG UK) made efforts to produce a representative sample of UK SEs, non-probability sampling strategies were used which reduces the potential for generalizability and statistical inference to the larger population of UK SEs. Recent UK government statistics suggests that this population is composed of 471 000 UK SEs, most of these (n=371 000) operating with no employees (Dept. BE&IS, UK Gov. 2017).

Nonetheless, descriptive statistics of key SE indicators (including percentage of organisations identified as registered charities, cooperatives, community businesses, social firms etc.) are comparable to responses of surveys conducted in 2011, 2013, and 2015, adding some external validity to our findings. Details of the survey methodology are available at Social Enterprise UK (2017:12). The questions used to derive the study variables appear in the online appendix.

In line with our top-level research question – that is, what factors predict SE social impact measurement – we looked for variables that would allow us to test the four hypotheses using logit regression analyses. Identifying the dependent variable was straightforward since the

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Note that these are not panel datasets and do not allow for tracking of SEs through time. Moreover, questions differ substantially between surveys preventing pooling of responses.
questionnaire directly asks SE respondents to rate his or her level of agreement on a scale of 1 (not at all) to 4 (to a large extent) with statement Q56: ‘To what extent does your organization measure its social impact?’ We note that while the survey does not explicitly define social impact, respondents are encouraged to differentiate between economic, environmental and social spheres of activity by the way the questions are structured and formulated. The social impact question (Q56) is located towards the end of the survey after the finance section. The survey does not refer to any particular social impact measurement systems or metrics. (See Table 1 for more details.)

Identifying the determinants of social impact measurement was more challenging since the questionnaire was not designed for this purpose. The first hypotheses (H1) investigate the extent to which SE context is linked to social impact measurement. Ideally, we would need information about proximity to metropolitan city centres but the survey only provides regional-level details of the geographical location of the SE although it includes ‘London’ are one of these regions (see online appendix). Accordingly, we constructed a dummy variable identifying whether the SE is located in or outside London (GEO-LONDON). We also constructed dummies that identified whether the SE belongs to a franchise or an organisation that replicates it’s management systems (FRANCHISE), and whether the SE recognizes itself as a cooperative (COOP). We also tested substitutes of these three dummies, including an urban-rural dummy (URBAN), dummies capturing alternative organisational forms such as whether the SE considered itself to be a social firm (SOCFIRM), a registered charity (RCHAR), or a community business (COMBUS), and dummies capturing its legal form, i.e. company limited by guarantee (CLBYG), company limited by shares (CLBYS), and community interest company limited by guarantee (CICCLG). We retained GEO-LONDON, FRANCHISE, and COOP for the regression analyses but used these substitutes in the robustness tests (see online appendix).
The second hypotheses (H2) investigate whether the way the SE envisages the nature of its social impact affects its decision to measure it. While all the SEs in our dataset participate in trading activities of various kinds, they differ in how they use these returns to achieve their social objectives. Accordingly, we constructed three dummies that capture whether the SE believes that its social impact lies primarily in the delivery of public services (SERVICES), the employment of disadvantaged individuals (EMPLOYMENT), or the creation of profits that are gifted to a separate cause or parent charity (PROFITS FOR SOCIAL GOODS).

The third and fourth hypotheses examine potential links between stakeholder pressures and impact measurement. To capture the influence of non-funding stakeholders (H3), we constructed three dummy variables from a questionnaire item that asked SEs to rate the extent to which community, staff, and beneficiaries were actively involved in SE decision-making (COMMUNITY, STAFF, BENEFICIARIES). Funding stakeholder pressures (H4) are derived from survey questions asking SEs to select their main source of income from a list of 13 choices. From these responses, we constructed two dummy variables entitled ‘TRADE/GRANTS PUBLIC SECTOR’ and ‘TRADE PRIVATE SECTOR’ although we also tested the impact of less frequently stated response categories including trade with general public, third sector organisations, donations, and membership or subscription fees.

For the controls, and in the absence of reliable financial data, we used proxies that capture access to (or lack of) financial resources. These included age, size, and scale of operation. More concretely, the survey questions that offered bandwidth response choices yielded lower non-response rates and we used these to generate dummies that captures whether the SE had been in operation 10 years or less (AGE - SE≤10 yrs) and similarly whether it had 10 or less employees (SIZE - SE≤10 emp). These cutoffs were determined by locating the median SE in the responses.
We also used 5 yrs or less and 5 employees or less versions of these dummies in the robustness checks (AGE - SE ≤ 5 yrs, SIZE - SE ≤ 5 emp). Scale of operation was also presented as 11 bandwidth choices (ranging from ‘your neighborhood’ to ‘internationally’) depending on the geographic reach of the SE’s operation. The most relevant indicator for our purposes was whether the SE’s operations were conducted on an international scale (SCALE – INTERNATIONAL).

More generally, using dichotomous variables throughout the analyses facilitates comparisons between odds ratios and result interpretation.

Descriptive statistics of the dependent and independent variables appear in Tables 1 and 2 while Table 1A in the online appendix provides point-biserial correlations (since the variables are dichotomous).

We used logit regression analyses to test our hypotheses after ruling out ordered logit regression. While ordered logit regression fully exploits the ordinal responses (see column (a) in Table 1), further testing using the Likelihood Ratio Test and the Brant Test demonstrated that the data did not support the parallel regression assumption required for this procedure (Long and Freese 2006). The OLS approach was also considered inferior, since the dependent variable only contains four ordinal response categories. Accordingly, we converted the responses to Q56 into a

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7 The ordered logit regression takes full advantage of the ordinal dependent variable. However, as it produces only one set of coefficients, the procedure implicitly assumes that the relationship between each pair of outcome groups is the same (the ‘parallel regressions assumption’). In other words, the procedure assumes that the coefficient when moving from MEASOCIMP=1 to MEASOCIMP=2 on the Likert Scale is the same as the coefficient that describes the transition from MEASOCIMP=2 to MEASOCIMP=3 and so on. If this is not the case, different models are required to describe the relationship between each pair of outcome groups.
dichotomous variable compatible with simple logit regression analysis by dividing responses symmetrically along the Likert Scale – i.e. scores 1 and 2 are assigned a ‘0’ and scores 3 and 4 are assigned a ‘1’ (see column (b)). Since there may be some ambiguity between scores 3 (‘to some extent’) and 2 (‘not very much’), we also tested versions of the dependent variable that would leave out responses in categories 3 (column (c)) and 2 (column (d)). The robustness tests were conducted using version (b) of the dependent variable as it allows the retention of all the data.

Explanatory variables were entered according to the four hypotheses in our framework (Figure 1) but alternative sequences were also tested. Ordered logit and OLS results are included for comparison. Since the variables used in the analyses are derived from the same survey instrument, we tested for Common Method Bias using Harman’s Single Factor Score. We also assess the extent of multicollinearity between variables.

A final parsimonious model was derived and reran on 36 subsamples of the data to assess the consistency of the results. More concretely, the subsamples in the online appendix compares odds ratios between SEs according to yrs of operation (Models 11-14), number of employees (Models 15-18), whether the SE is located in a urban or rural area (URBAN - Models 19 and 20) and investments in employee training (TRAIN - Models 21 and 22). Models 23 to 34 compare odds ratios across SEs with different organisational and legal forms as discussed above. Models 35-46 compares odds ratios across SEs with different organisational capabilities including people management (PEOPMGT), developing and implementing a business plan and strategy (BUDP&S), developing and introducing new products or services (NEWPROD), making effective use of available technology (TECH), financial management (FINMGT), marketing, branding and PR (MARKT), and reacting to regulation and tax issues (REGTX).
Results

Table 3 presents regression results conducted on the entire dataset. Model 1 includes only the control variables. Of the three, the odds ratio on the SCALE-INTERNATIONAL variable is above one throughout all the analyses and statistically significant suggesting a positive relationship with SE propensity to measure social impact. The odds ratio on AGE is also above one but its significance varies considerably across the subsamples while the odds ratio on the SIZE variable is only marginally significant on two occasions (Models 20 and 41).

The three variables capturing SE context (H1) are added in Model 2. All the odds ratios are significant at the level \( p=0.05 \) or less. The odds ratios for GEO-LONDON and FRANCHISE are above one, whereas the odds ratio on COOP is negative suggesting an inverse relationship with social impact measurement. Our results align with H1 but adding the additional variables and testing them across subsamples help determine their strength. None of the odds ratio of the variables relating to the nature of SE impact (H2) are significant (Model 3). This is a surprising result since the ease with which social impact can be measured across these three sets of activities will vary considerably. Adding these variables do not change the strength nor significance of the odds ratio relating to H1 variables.

H3 and H4 posit that SEs may be more willing to measure their social impact if pressured by their stakeholders (Models 4 and 5). We find that SEs who involve community and staff members in their decision-making are more likely to measure social impact - i.e. the odds ratios are greater than one and significant - but the odds ratios on the involvement of beneficiaries are not significant. As for funding stakeholders (H4), only the odds ratio on the TRADE/GRANTS PUBLIC SECTOR indicator are statistically significant. These are greater than one confirming the hypothesis that SEs whose main income is from government are more likely to measure social
impact. We tested the order in which the variables were entered and found no significant differences in the results.

Model 6 presents a parsimonious version of the model where non-significant variables are removed. This reduced specification is used for further testing. Models 7 and 8 examine the sensitivity of the results to different versions of the dependent variable and the treatment of the ‘middle’ response categories (see Table 1). The results remain qualitatively unchanged providing support for our use of a dependent variable that identifies SEs with scores 3 and 4 as ‘measuring social impact’ and those with scores 1 and 2 are ‘not measuring social impact’. Models 9 and 10 display the results from the ordered logit and OLS regressions as benchmarks. Multicollinearity was ruled out since the highest VIF we encountered was 1.050. Using the Harman’s Single Factor Score approach, we found that none of the factors was in excess of 50% suggesting that CMB is unlikely to affect our results (Podsakoff, Mackenzie, Lee and Podsakoff 2003).

____________________________ Insert Table 3 here __________________________________

We subsequently tested the logit regression results across 36 different subsamples of the data to assess their robustness (see online appendix, Tables A2-A4). We found surprising consistency throughout the subsamples with no observed qualitative changes in the nature of the odds ratios (i.e. those greater than one remained so and vice versa) although there is weaker or absence of significance with smaller samples sizes (i.e. standard errors tend to be greater in smaller samples, reducing the statistical significance of the odds ratios). We note in particular that the odds ratios associated with FRANCHISE and COMMUNITY are consistently above one while those on the COOP variable are always below one and statistically significant throughout almost all of
the 44 models. The strength of these results is reflected by the (++) and (--) symbols in the results summary (Figure 2). The sign of the odd ratios on the other variables (GEO-LONDON, STAFF, and TRADE/GRANTS PUBLIC SECTOR) remain above one indicating positive relationships with the SE propensity to measure social impact but we observe more instances of non-significance particularly with smaller sample sizes. These relationships are denoted (+) in Figure 2 while the variables that were not significant in the analyses are crossed out.

Discussion

Our results point to the context in which a particular SE operates being an important predictor of the likelihood that they measure social impact. SEs located in London may be more prone to normative isomorphic pressures as they are embedded in philanthropic networks that are more familiar, and converted to, NPM-style measurement schemes (McCulloch et al 2012; Mohan 2012; Millar and Hall 2013; Ebrahim and Rangan 2014). Similarly, franchised SEs will typically have greater access to expert advice and more pressures to adopt tried-and-tested management systems that will facilitate, amongst other things, the measurement of social impact (Dees et al 2004; VanSandt et al 2009; Bloom and Chatterji 2009; Lyon and Fernandez 2012). The results also indicate that SEs who identify themselves as COOPs are less likely to measure social impact, possibly reflecting COOP traditions of self-management and resistance to outside interference (Rothschild 2009; Doherty et al 2014; Defourny and Nyssens 2010; Ebrahim et al 2014). We therefore find support for our claim that SEs who have easier access to knowledge and professional resources are more likely to measure social impact (H1a and H1b) while the COOP form seems to dampen conformance with dominant sector accountability norms and pressures (H1c).
The analysis did not support the hypotheses linking social impact measurement to differences in the ‘measurability’ SE activities (H2) (Doherty et al 2014; Millar and Hall 2013; Bagnoli and Megali 2011; Maas and Greco 2017). This is surprising. Our findings suggest that either the complexity of measurement is not a significant factor in SE decisions and/or that our typology is not sufficiently refined to capture differences in the ‘difficulty’ of measurement processes across settings. It may also be the case that our dependent variable is too rudimental and that a richer measure that accounts for the intensity of the measurement effort and/or the use of specific metrics would produce different results. This highlights the need for further sector-wide data collection and analyses (Saebi Foss and Linder 2019).

The literature suggests that social impact measurement can enhance SE institutional legitimacy by facilitating communications and accountability reporting to stakeholders and creating the reference frames that mobilise their support (e.g. Bagnoli and Megali 2011; Nicholls 2010a, 2010b; Lee and Huang 2018). While this literature highlights differences in pressures exerted on SEs by funders versus other stakeholder groups (O’Dwyer and Unerman 2007; Arvidson and Lyon 2014; Barman 2007; Millar and Hall 2013), relatively little is known about differences in how non-funding stakeholders such as community members, staff, and beneficiaries affect SEs processes. Our findings that community and staff engagement in decision-making are linked to the utilization of social impact measures gives credence to these claims (H3a, H3b) although it is not clear why beneficiaries do not impact on the dependent variable (H3c). One possibility is that social impact measures are not effective in communications with this particular stakeholder group and/or that beneficiaries express resistance or indifference to such development (Arvidson and Lyon 2014). Our results thus partially support the hypothesis that SEs practicing multi-stakeholder governance may be more willing to measure their social impact or conversely,
that measurement facilitates multi-stakeholder governance (H3). These findings emphasize our limited understanding of SE stakeholder salience to the organisation but highlight an interesting avenue for future research.

The data and analyses also support our hypothesis that social impact measurement is associated with public sector trade and grants (H4a) but does not support its private sector counterpart (H4b). These findings corroborate insights from the literature suggesting that seeking funding from public sector entities, either through grants, trade or contracts, will ‘coerce’ SEs into adopting NPM-style reporting and accountability mechanisms (Arvidson and Lyon 2014; Nicholls 2010b; Gibbon and Dey 2011; Cunningham et al, 2014). It may also be the case that SEs who have the resources to measure social impact are more likely to seek such funds (Liston-Heyes et al 2017; Julnes and Holzer 2001).

Regardless of causality, our results underline the scale and scope of public sector bonds to the UK SE sector.

**Conclusion**

Social impact measures have become important determinants of SE legitimacy and government funding in the UK and elsewhere. Yet not all SEs are opting to measure their social impact despite strong institutional pressures to do so. Some are isolated from these pressures while others believe that conformance is either too costly or damaging to the SE mission. In this study we attempt to document the patterns of social impact measurement utilisation in the UK SE sector.

Our study meets three objectives. First, it reviews the literature to identify potential sources of isomorphic pressures exerted on SEs by the public sector in the adoption of social impact measurement. The resulting narrative explains how and why the sector became under increasing
pressure to conform to public sector norms of comparability, audit, and value-for-money. It highlights in particular how grants and contract allocations, reporting processes, training schemes, and consultants may be incentivising SEs into measuring their social impact. Our review also suggests that the absence of an established governance framework catering to the specificity of the hybrid form may have contributed to the spread of NPM-style indicators.

Secondly, our study presents a model of social impact measurement adoption that exploits these stylised facts. The model suggests that some of the pressures exerted on SEs will be embedded in the context in which they operate (H1) and that measurement scheme adoption may vary according to the ‘measurability’ of the impact produced by SEs (H2). It also postulates that social impact measurement will be linked to SE engagement with non-funding stakeholders (beneficiaries, staff, and community) (H3), and funding stakeholders (public sector, private sector) (H4). The resulting framework provides a tool to help differentiate between SEs who measure social impact and those who do not.

Its third objective was to test the model on a recent dataset of UK SEs. Our logit regression results show that SEs that are based in London, franchised, and who have the public sector as their main income source are more likely to measure their social impact while SEs that identify themselves as COOP are generally less likely to do so. We also find that SEs who involve community stakeholders - and to a weaker extent their staff - in their decision-making are more likely to measure social impact. Somewhat surprisingly, the nature of SE impact carried no explanatory power in the adoption decision.

Our results provide a number of interesting insights into SE impact measurement. Firstly, they indicate that SEs are more likely to measure their social impact if governmental agencies are their main source of funding. This supports the notion that the public sector has been instrumental
in shaping the UK’s SE sector by championing the potential of SEs as a novel way of delivering public services, by transferring public service delivery responsibilities to the sector in the aftermath of the financial crisis, and by endorsing NPM-inspired accountability discourse surrounding SE performance. Since SEs are often depicted as combining elements from the private and charitable sectors, our results emphasise the influence the public sector is exerting on these organisations as hypothesised and corroborated by existing case studies. Secondly, the link between context factors (e.g. access to networks and location) and measurement decisions emphasises the strength of institutional pressures while highlighting the importance of sector associations, training, professionalization, and labour markets. Thirdly, our results show that SEs that engage their community and staff stakeholders in their decision-making processes are more likely to measure social impact but this is not the case with SE beneficiaries. This suggests potential differences in the relative benefits of social measures across stakeholder groups, of relevance to those advocating their use in multi-stakeholder governance models. Finally, our study underscores the dearth of sector-wide information available on SEs and the need for further investment in systematic and regular data collection exercises.

In learning from the results it is important to bear in mind some caveats. The dataset, while the largest and most up-to-date for the UK sector, was constructed using non-probability sampling strategies. Since this is not a random sample, care is needed in extrapolating results to the wider population. Our results describe empirical patterns found in this subset of the population. Second, our regressions help identify the characteristics of SEs who measure social impact versus those who do not. While we postulate on the mechanisms underlying these relationships, our analyses do not establish causality. Thirdly, our dependent variable does not assess the SE’s interpretation of social impact, the ‘type’ of impact measurement scheme that is used, or the extent with which
SE managers utilise it. This implies that what constitutes social impact is likely to vary considerably across SEs. Moreover, some SEs may be measuring social impact but only in a symbolic way, without fully exploiting its potential or endorsing its discourse. Our analyses do not capture these important nuances. Finally, our study identifies London as a major metropolitan area. We recognise that the UK has other metropolitan cities (e.g. Birmingham, Manchester, and Leeds) that may also provide access to social capital, experienced professionals, and the networks that can encourage and facilitate social impact measurement. Unfortunately, the dataset classifies SEs by regions except for London-based organisations. While additional tests indicate that an urban-rural divide does not capture these differences in environment and that even amongst urban-based SEs, there appears to be a London-specific effect, additional research is needed to explore this issue to its fullest extent.

Barman (2007:112) concludes her historical account by stating that we should ‘… view the extent of nonprofits’ need to quantity as the inverse of the size and scope of the central government’. Our study provides empirical support for this narrative and complements it by providing preliminary results on who is quantifying and possibly why.
References


Table 1 Dependent Variable (MEASOCIMP)

Responses to Q56 ‘To what extent does your organisation measure its social impact?’

<table>
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<tr>
<th>Values</th>
<th>Frequency</th>
<th>Frequency</th>
<th>Values</th>
<th>Frequency</th>
<th>Frequency</th>
<th>Values</th>
<th>Frequency</th>
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</thead>
<tbody>
<tr>
<td>4 - A large extent</td>
<td>400</td>
<td>1</td>
<td>832</td>
<td>1</td>
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<td>832</td>
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<td>1</td>
<td>removed</td>
<td>432</td>
<td>1</td>
<td>832</td>
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<td>2 - Not very much</td>
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<td>135</td>
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<td>1 - Not at all</td>
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<tr>
<td>Total</td>
<td>1060</td>
<td>Total</td>
<td>1060</td>
<td>Total</td>
<td>628</td>
<td>Total</td>
<td>925</td>
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Table 2 Independent Variables

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<tr>
<th>Variables</th>
<th>Frequency</th>
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<tr>
<td><strong>SE Context (H1)</strong></td>
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<tr>
<td>GEO - LONDON</td>
<td>248</td>
<td>23%</td>
</tr>
<tr>
<td>FRANCHISE</td>
<td>68</td>
<td>6.3%</td>
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<tr>
<td>COOP</td>
<td>227</td>
<td>20.9%</td>
</tr>
<tr>
<td><strong>Nature of SE impact (H2)</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td>SERVICES</td>
<td>385</td>
<td>35.5%</td>
</tr>
<tr>
<td>EMPLOYMENT</td>
<td>218</td>
<td>20.1%</td>
</tr>
<tr>
<td>PROFITS FOR SOCIAL GOODS</td>
<td>254</td>
<td>23.5%</td>
</tr>
<tr>
<td><strong>SE Non-Funding Stakeholders (H3)</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td>COMMUNITY</td>
<td>591</td>
<td>54.5%</td>
</tr>
<tr>
<td>STAFF</td>
<td>956</td>
<td>88.2%</td>
</tr>
<tr>
<td>BENEFICIARIES</td>
<td>768</td>
<td>70.8%</td>
</tr>
<tr>
<td><strong>SE Funding Stakeholders (H4)</strong></td>
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<td></td>
</tr>
<tr>
<td>TRADE/GRANTS PUBLIC SECTOR</td>
<td>290</td>
<td>26.8%</td>
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<tr>
<td>TRADE PRIVATE SECTOR</td>
<td>159</td>
<td>14.7%</td>
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<tr>
<td><strong>Controls</strong></td>
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<td></td>
</tr>
<tr>
<td>AGE (SE &lt; 10 yrs)</td>
<td>661</td>
<td>61.0%</td>
</tr>
<tr>
<td>SIZE (SE &lt; 10 emp)</td>
<td>740</td>
<td>68.3%</td>
</tr>
<tr>
<td>SCALE – INTERNATIONAL</td>
<td>119</td>
<td>11%</td>
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</table>
Table 3 Logit, Ordered Logit, and OLS Regression Analyses

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<tbody>
<tr>
<td>Constant</td>
<td>3.2882*** (.4858)</td>
<td>3.1905*** (.5236)</td>
<td>2.8353*** (.5657)</td>
<td>1.1594 (.3431)</td>
<td>1.0839 (.3295)</td>
<td>1.0808 (.2986)</td>
<td>.4422** (.1451)</td>
<td>2.2136*(.8679)</td>
<td>N.A.</td>
<td>.5741*** (.0460)</td>
</tr>
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</table>

H2 Nature of SE Impact

<p>| | | | | | | | | | | |</p>
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</tr>
</thead>
<tbody>
<tr>
<td>SERVICES</td>
<td>1.4394* (.2787)</td>
<td>1.2541 (.2505)</td>
<td>1.1592 (.2364)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>EMPLOYMENT</td>
<td>.8299 (.1730)</td>
<td>.8893 (.1915)</td>
<td>.8903 (.1925)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>PROFITS FOR SOCIAL GOODS</td>
<td>1.070 (.2827)</td>
<td>1.0119 (.2737)</td>
<td>1.0169 (.2775)</td>
<td></td>
<td></td>
<td></td>
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</tr>
</tbody>
</table>

H4 Funding Stakeholders

|                      |                      |                      |                      |                      |                      |                      |                      |                      |                      |                      |
| TRADE/GRANTS PUB. SECTOR | 1.4255* (.2995)    | 1.5430* (.3072)      | 1.4865* (.3364)      | 1.8928*** (.6035)    | .1338 (.1344)         | .0593* (.2818)         |                      |                      |                      |                      |
| TRADE/GRANTS PRIV. SECTOR | .8563 (.1905)     |                      |                      |                      |                      |                      |                      |                      |                      |                      |

Controls

|                      |                      |                      |                      |                      |                      |                      |                      |                      |                      |                      |
| AGE (SE < 10 yrs)    | 1.5773*** (.2508)    | 1.3814* (.2265)      | 1.3562* (.2245)      | 1.2865 (.2181)       | 1.2886 (.2191)       | 1.2811 (.2159)       | 1.4837* (.2952)       | 1.6125** (.3888)       | 1.2639** (.1257)      | .0409* (.2626)        |
| SIZE (SE < 10 employees) | .7213* (.1242)     | .9245 (.1661)        | .9682 (.1759)        | .9750 (.1834)        | 1.0291 (.1976)       | 1.0332 (.1960)       | 1.0144 (.2246)       | .8864 (.2580)          | -.0324 (.1356)        | -.00075 (.2825)       |
| SCALE INTERNATIONAL | 1.9680* (.5688)     | 1.6370* (.4831)      | 1.7957* (.5436)      | 1.8499* (.5621)      | 1.9268* (.5787)      | 1.8669* (.6368)      | 2.0100* (.6838)       | 1.4904 (.6389)         | .3859* (.1900)        | .0744* (.5387)        |

N observations 1060 1060 1060 1060 1060 1060 1060 1060 1060 1060 1060
-Log likelihood 543.8674 519.9539 516.2780 492.73967 490.61467 491.67458 547.92941 262.6431 1225.3299
Root MSE 39051
LR chi2 16.00 63.82 71.18 118.25 122.50 120.38 127.02 78.31 113.53
F 13.72
Pseudo R .0145 .0578 .0645 .1071 .1110 .1091 .1544 .1297 .0443
R/ Adj R .117.10
Prob > chi2 .0011 .0000 .0000 .0000 .0000 .0000 .0000 .0000 .0000
Prob > F .0000

***, **, * denote p values <.001, <.01, <.05 and <.10 respectively.

a Parsimonious regression.
b Versions ‘c’ and ‘d’ of the dependent variable – please see Table 1.
d Ordered Logit and OLS regressions results respectively.
Figure 1: To Measure or Not to Measure?

Does the SE measure its social impact?

- **SE Context**: GEOG – LONDON, FRANCHISE, COOP
- **Non-Funding Stakeholders**: COMMUNITY, STAFF, BENEFICIARIES
- **Nature of SE Impact**: SERVICES, EMPLOYMENT, PROFITS FOR SOCIAL GOODS
- **Funding Stakeholders**: TRADE/GRANTS PUBLIC SECTOR, TRADE/GRANTS PRIVATE SECTOR

**Hypotheses**:
- **H1**: The SE context.
- **H2**: The nature of SE impact.
- **H3**: Non-funding stakeholders.
- **H4**: Funding stakeholders.

**Controls**: SE age; SE size; SE operating scale.

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Figure 2: Results Synthesis

Does the SE measure its social impact?

- **SE Context**: GEOG – LONDON ‘+++’, FRANCHISE ‘+’, COOP ‘-’
- **Non-Funding Stakeholders**: COMMUNITY ‘+++’, STAFF ‘++’, BENEFICIARIES
- **Nature of SE Impact**: SERVICES, EMPLOYMENT, PROFITS FOR SOCIAL GOODS
- **Funding Stakeholders**: TRADE/GRANTS PUBLIC SECTOR ‘+++’, TRADE/GRANTS PRIVATE SECTOR ‘-’

**Legend**
- ‘+++’: positive and statistically significant throughout the analyses.
- ‘++’: negative and statistically significant throughout the analyses.
- ‘+’: positive and statistically significant throughout many analyses.
- ‘-’: positive and statistically significant throughout many analyses.
- Strikethrough: not statistically significant.